

- 6.78 Let Y_1 be the point on the one-mile stretch chosen for sentry 1 and let Y_2 be the point chosen for sentry 2. It is given that $f_1(y_1) = 1$ for $0 \leq y_1 \leq 1$ and $f_2(y_2) = 1$ for $0 \leq y_2 \leq 1$. Since Y_1 and Y_2 are independent, their joint distribution is
- $$f(y_1, y_2) = f_1(y_1)f_2(y_2) = 1 \quad 0 \leq y_1 \leq 1, 0 \leq y_2 \leq 1$$

The distance between the two sentries, which is the variable of interest, can be represented by $|Y_1 - Y_2|$. Let us consider first the joint distribution of $u = Y_1 - Y_2$ and Y_2 . Letting $Y_2 = y_2$, we have $Y_1 = U + y_2$ and $|\frac{dy_1}{du}| = 1$. Hence $g(y_2, u) = 1$. The region on which the density is positive is shown in Figure 6.10.

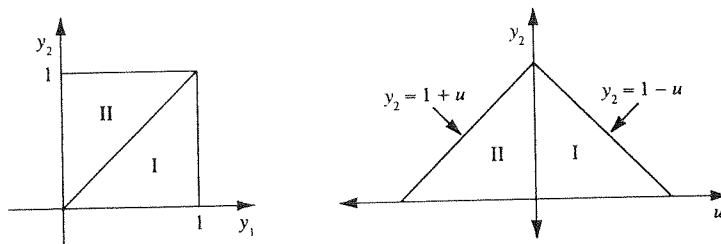


Figure 6.10

$$g(u) = \int_0^{1+u} dy_2 = 1 + u \quad \text{for} \quad -1 \leq u \leq 0$$

$$g(u) = \int_0^{1-u} dy_2 = 1 - u \quad \text{for} \quad 0 \leq u \leq 1.$$

Now the probability of interest is

$$\begin{aligned} P(|Y_1 - Y_2| < \frac{1}{2}) &= P(|U| < \frac{1}{2}) = P(-\frac{1}{2} \leq U \leq \frac{1}{2}) \\ &= \int_{-1/2}^0 (1 + u) du + \int_0^{1/2} (1 - u) du \\ &= \frac{3}{4} \end{aligned}$$

- 7.2** a. $P(|\bar{Y} - \mu| \leq .3) = P[-.3 \leq (\bar{Y} - \mu) \leq .3] = P\left(\frac{-.3}{\left(\frac{\sigma}{\sqrt{n}}\right)} \leq \frac{Y - \mu}{\left(\frac{\sigma}{\sqrt{n}}\right)} \leq \frac{.3}{\left(\frac{\sigma}{\sqrt{n}}\right)}\right)$
 $= P(-.3\sqrt{n}/2 \leq z \leq .3\sqrt{n}/2) = P(-.15\sqrt{n} \leq z \leq .15\sqrt{n})$
 $= 1 - 2P(z > .15\sqrt{n}) = 1 - 2P(z > .45) = 1 - 2(.3264) = .3472$. This probability is less than that found in example 7.1 (as σ has increased).
- b. The probability in question is $1 - 2P(z > .15\sqrt{n})$ which is $1 - 2P(z > .75) = .5468$ for $n = 25$, $.6318$ for $n = 36$, $.7062$ for $n = 49$, and $.7698$ for $n = 64$.
- c. The probabilities increase as n increases.
- d. As we increase σ the variability of \bar{Y} increases. Therefore the probabilities for exercise 7.2 part b should be smaller than those for exercise 7.1 part b. (which is exactly what we see).

- 7.10** a. Since a χ^2 random variable is defined as a gamma random variable with $\alpha = \frac{\nu}{2}$ and $\beta = 2$, the expected value and variance are
 $E(U) = \alpha\beta = \nu$ and $V(U) = \alpha\beta^2 = 2\nu$
- b. Using Theorem 7.3, we see that the quantity $\frac{(n-1)S^2}{\sigma^2}$ has a χ^2 distribution with $\nu = n - 1$. Hence, from part a,
 $E\left(\frac{(n-1)S^2}{\sigma^2}\right) = n - 1$ or $\left(\frac{n-1}{\sigma^2}\right)E(S^2) = n - 1$ or $E(S^2) = \sigma^2$
 Similarly,
 $V\left(\frac{(n-1)S^2}{\sigma^2}\right) = 2(n - 1)$ or $\left[\frac{(n-1)^2}{\sigma^4}\right]V(S^2) = 2(n - 1)$ or $V(S^2) = \frac{2\sigma^4}{n-1}$

7.12 Similar to Exercise 7.11. Since, from Definition 7.2,

$$P(g_1 \leq (\bar{Y} - \mu) \leq g_2) = P\left(\frac{g_1}{\left(\frac{\sigma}{\sqrt{n}}\right)} \leq t \leq \frac{g_2}{\left(\frac{\sigma}{\sqrt{n}}\right)}\right)$$

the necessary probability will be obtained if we take

$$\frac{g_1}{\left(\frac{\sigma}{\sqrt{n}}\right)} = -t_{.05} \quad \text{and} \quad \frac{g_2}{\left(\frac{\sigma}{\sqrt{n}}\right)} = t_{.05}$$

Indexing $n - 1 = 8$ degrees of freedom, we have $t_{.05} = 1.86$, $g_1 = \left(\frac{-1.86}{3}\right)S$, and $g_2 = \left(\frac{1.86}{3}\right)S$.